Bayesian reconstruction of the Cosmological Large Scale Structure

Francisco-Shu Kitaura

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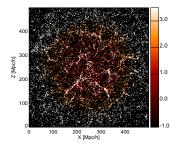






Introduction

- the cosmological large-scale structure encodes a wealth of information about the evolution and origin of the Universe
- the data are plagged by many observational effects (mask, selection function, bias ...)
- statistical treatment is necessary
- compare observations with theory
- study structure formation



 $\begin{array}{l} \longrightarrow \mbox{ matter field } \delta_{\rm M}? \\ \longrightarrow \mbox{ pec. vel. field } {\bf v}? \\ \longrightarrow \mbox{ grav. pot. } \Phi_{\rm Grav}? \end{array}$

Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

(1)

Bayes theorem: the posterior

$$\mathcal{P}(\mathbf{s}|\mathbf{d},\mathbf{p}) = rac{\mathcal{P}(\mathbf{s}|\mathbf{p})\mathcal{P}(\mathbf{d}|\mathbf{s},\mathbf{p})}{\int \mathrm{d}\mathbf{s}\,\mathcal{P}(\mathbf{s}|\mathbf{p})\mathcal{P}(\mathbf{d}|\mathbf{s},\mathbf{p})},$$

Posterior=prior×likelihood/evidence

 \longrightarrow clear representation of the assumptions \longrightarrow can be easily extended to nonlinear cases

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Bayesian inference steps

Definition of the prior: knowledge of the underlying signal

- Definition of the likelihood: nature of the observed data
- Linking the prior to the likelihood: link signal to the data
- Bayes theorem: the posterior
- Maximization of the posterior: MAP
- Sampling the posterior: MCMC

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Physical motivation: model | FK & Raul Angulo in prep

Lagrangian perturbation theory: displacement field

$$\mathbf{x} = \mathbf{q} + \mathbf{\Psi} \,. \tag{2}$$

Mass conservation:

$$\rho(\mathbf{x}, t) \mathrm{d}\mathbf{x} = \langle \rho(t_i) \rangle \mathrm{d}\mathbf{q} \,.$$
(3)

The inverse of the Jacobean leads to the overdensity field:

$$1 + \delta(\mathbf{x}(\mathbf{q}, t)) = \mathbf{J}(\mathbf{q}, t)^{-1}, \qquad (4)$$

with

$$\mathbf{J}(\mathbf{q},t) = \left|\frac{\partial \mathbf{x}}{\partial \mathbf{q}}\right| \,.$$

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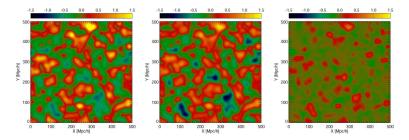
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$$\mathbf{J}(\mathbf{q},t) = \left| \frac{\partial \mathbf{x}}{\partial \mathbf{q}} \right| \,. \tag{5}$$

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Testing model I with MR

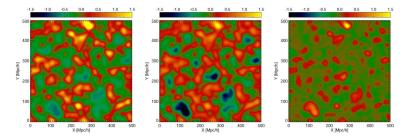
Forward relation from z = 127 to z = 0:



Testing model I with MR

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Inverse relation from z = 0 to z = 127:

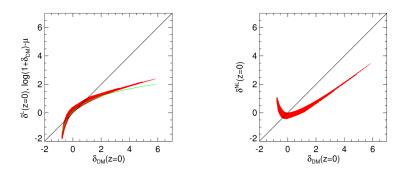


Definition of the prior Definition of the likelihood Link between the prior and the

Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Testing model I and II with MR FK & Raul Angulo in prep

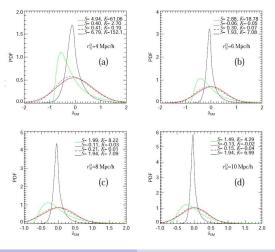
Cell-to-cell comparison:



Definition of the prior

Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Testing model I and II with MR: matter statistics

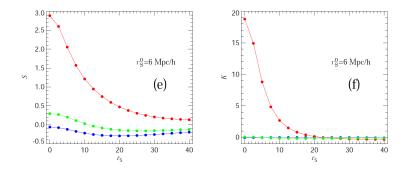


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Bayesian reconstruction

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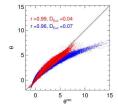
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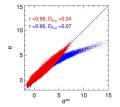


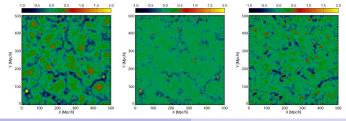
Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Testing model I and II with MR: peculiar motions

FK, Raul Angulo, Yehuda Hoffman & Stefan Gottloeber in prep



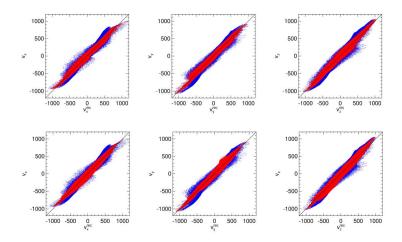




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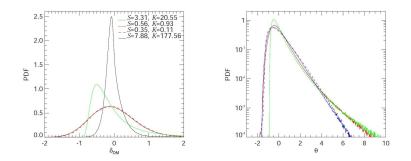
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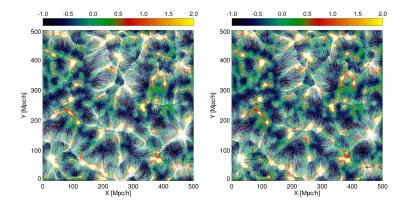


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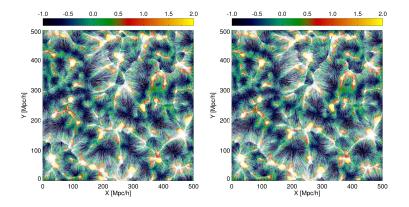
Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling



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Non-Gaussian expansions

- relax the Gaussian-Lognormal assumption
- skewed Gaussian Edgeworth expansion (univariate case: Juszkiewicz, Bouchet & Colombi 93)
- skewed lognormal model with the 1D Edgeworth exansion (univariate case: Colombi 94)

Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Non-Gaussian expansions: multivariate case FK 2010

Let me introduce here the multivariate case:

$$\Phi_i \equiv \ln \rho_i - \langle \ln \rho \rangle = s_i - \mu_i, \quad \nu_i \equiv \sum_j S_{ij}^{-1/2} \Phi_i , \quad (6)$$

$$s_i = \ln(
ho_i/\langle
ho
angle) = \ln(1+\delta_{\mathrm{M}i})$$

Multidimensional Edgeworth expansion

$$P(\mathbf{\Phi}) = G(\nu) \Big[1 + \frac{1}{3!} \sum_{i'j'k'} \langle \Phi_{i'} \Phi_{j'} \Phi_{k'} \rangle \sum_{ijk} S_{ii'}^{-1/2} S_{jj'}^{-1/2} S_{kk'}^{-1/2} h_{ijk}(\nu) + \frac{1}{4!} \sum_{i'j'k'l'} \langle \Phi_{i'} \Phi_{j'} \Phi_{k'} \Phi_{l'} \rangle \sum_{ijkl} S_{ii'}^{-1/2} S_{jj'}^{-1/2} S_{kk'}^{-1/2} S_{ll'}^{-1/2} h_{ijkl}(\nu) + \dots \Big],$$
(7)

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Non-Gaussian expansions: multivariate case FK 2010

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Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Lognormal model

• Oth order: lognormal model $(G(\nu) \rightarrow \mathcal{P}(\delta_{\mathrm{M}}|\mathbf{S}))$ $\langle \ln \rho_i \rangle = \ln \langle \rho \rangle + \mu_i$

$$\mathcal{P}(\boldsymbol{\delta}_{\mathrm{M}}|\mathbf{S}) = \frac{1}{\sqrt{(2\pi)^{N_{\mathrm{cells}}} \mathrm{det}(\mathbf{S})}} \prod_{k} \frac{1}{1 + \delta_{\mathrm{M}k}}$$

$$\times \exp\left(-\frac{1}{2} \sum_{ij} \left(\ln(1 + \delta_{\mathrm{M}i}) - \mu_{i}\right) S_{ij}^{-1} \left(\ln(1 + \delta_{\mathrm{M}j}) - \mu_{j}\right)\right),$$
(8)

multidimensional implementation for matter field reconstructions: FK, Jasche & Metcalf 2009; applied to SDSS DR7: Jasche, FK, Li & Ensslin 2010 when $\delta_M \ll 1 \rightarrow$ Gauss distribution

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Definition of the likelihood

- Discreteness of the galaxy distribution: shot noise
- Galaxy selection function in a magnitude limited survey
- Link between underlying matter field and the galaxy field
- Discrete Press-Schechter: Borel distribution (Epstein 1983)
 Gravitothermal dynamics (Saslaw 1986, Itoh et al 1988, Sheth 1995)
- Non-Poissonian distribution (FK in prep)

$$\mathcal{P}(\{N_{k}^{g}\}|\{\lambda_{k}\},\mathbf{Q}) = \prod_{k} \frac{\sum_{j} (\delta_{k,j}^{K} - Q_{k,j})\lambda_{j}}{N_{k}^{g}!}$$
$$\times \left(\sum_{l} (\delta_{k,l}^{K} - Q_{k,l})\lambda_{l} + \sum_{m} Q_{k,m} N_{m}^{g}\right)^{N_{k}^{g}-1}$$
$$\times \exp\left(-\sum_{n} (\delta_{k,n}^{K} - Q_{k,n})\lambda_{n} - \sum_{o} Q_{k,o} N_{o}^{g}\right)$$
(9)

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Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Define the likelihood: nature of the observable

For a sparse sample: Poisson limit

$$\mathcal{P}(\{N_k^{\mathrm{g}}\}|\{\lambda_k\}) = \prod_k \frac{\lambda_k^{N_k^{\mathrm{g}}} \exp\left(-\lambda_k\right)}{N_k^{\mathrm{g}}!} \tag{10}$$

full treatment: FK & Ensslin 2008; FK, Jasche & Metcalf 2009

Observation process: radial selection function and sky mask binomial process: either we see the galaxy or not

$$\lambda_i \equiv w_i \lambda_i' \tag{11}$$

treatment proposed in FK, Jasche, Li, Ensslin, Metcalf, Wandelt, Lemson & White 2009

• the correlation is encoded in the underlying density field in λ_k

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Link between the prior and the likelihood

galaxy bias

$$\delta_{\mathrm{g}i} = B(\boldsymbol{\delta}_{\mathrm{M}})_i, \qquad (12)$$

Fry & Gaztanaga 93 (generalized)

$$\delta_{gi} = \sum_{j} B_{ij}^{1} \delta_{Mj} + \delta_{Mi} \sum_{j} B_{ij}^{2} \delta_{Mj} + \dots, \qquad (13)$$

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Response operator

Response operator

$$\lambda_{k} = \lambda_{k}(\boldsymbol{\delta}_{\mathrm{M}}) = R(\boldsymbol{\delta}_{\mathrm{g}}(\boldsymbol{\delta}_{\mathrm{M}}))_{k}, \qquad (14)$$

$$\lambda_k = w_k \bar{N} (1 + B(\delta_{\mathrm{M}})_k), \qquad (15)$$

$$\lambda_k = w_k \bar{N} (1 + b \delta_{\mathrm{M}k}), \qquad (16)$$

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(17)

Bayes theorem: the posterior

$$\mathcal{P}(\boldsymbol{\delta}_{\mathrm{M}}|\mathbf{\mathsf{N}},\mathbf{\mathsf{S}}) = rac{\mathcal{P}(\boldsymbol{\delta}_{\mathrm{M}}|\mathbf{\mathsf{S}})\mathcal{P}(\mathbf{\mathsf{N}}|\boldsymbol{\lambda}(\boldsymbol{\delta}_{\mathrm{M}}))}{\int\mathrm{d}\boldsymbol{\delta}_{\mathrm{M}}\,\mathcal{P}(\boldsymbol{\delta}_{\mathrm{M}}|\mathbf{\mathsf{S}})\mathcal{P}(\mathbf{\mathsf{N}}_{k}|\boldsymbol{\lambda}(\boldsymbol{\delta}_{\mathrm{M}}))},$$

Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Bayes theorem: the posterior

$$\mathcal{P}(\boldsymbol{\delta}_{\mathrm{M}}|\mathbf{N},\mathbf{S}) \propto \prod_{i} \frac{1}{1+\delta_{\mathrm{M}i}} \exp\left(-\frac{1}{2}\sum_{ij}\left(\ln\left(1+\delta_{\mathrm{M}i}\right)-\mu_{i}\right)S_{ij}^{-1}\left(\ln\left(1+\delta_{\mathrm{M}j}\right)-\mu_{j}\right)\right) \times \left[1+\frac{1}{3!}\sum_{i'j'k'}\left\langle\Phi_{i'}\Phi_{j'}\Phi_{k'}\right\rangle\sum_{ijk}S_{ii'}^{-1/2}S_{jj'}^{-1/2}S_{kk'}^{-1/2}h_{ijk}(\boldsymbol{\nu}) + \frac{1}{4!}\sum_{i'j'k'l'}\left\langle\Phi_{i'}\Phi_{j'}\Phi_{k'}\Phi_{l'}\right\rangle\sum_{ijkl}S_{ii'}^{-1/2}S_{jj'}^{-1/2}S_{kk'}^{-1/2}S_{ll'}^{-1/2}h_{ijkl}(\boldsymbol{\nu}) + \dots\right] \times \prod_{k}\frac{\left(w_{k}\bar{N}(1+b\delta_{\mathrm{M}k})\right)^{N_{k}}\exp\left(-w_{k}\bar{N}(1+b\delta_{\mathrm{M}k})\right)}{N_{k}!},$$
(18)

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Bayes theorem: the posterior

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Maximum a posteriori

• Let us define the energy $E(\mathbf{s})$

$$E(\mathbf{s}) \equiv -\ln\left(\mathcal{P}\left(\mathbf{s}|\mathbf{N},\mathbf{S}\right)\right),\tag{19}$$

MAP

$$\frac{\partial E(\mathbf{s})}{\partial s_l} = 0, \tag{20}$$

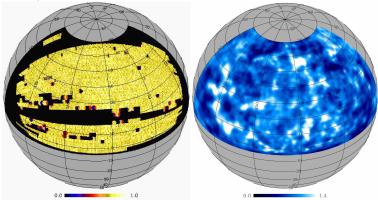
 Krylov conjugate gradient schemes (FK & Ensslin 2008; Jasche, FK, Wandelt, Ensslin 2009; FK, Jasche, & Metcalf 2009)

$$s_i^{j+1} = s_i^j - \sum_k T_{ik} \frac{\partial E(\mathbf{s})}{\partial s_k}, \qquad (21)$$

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Results: Wiener filter reconstruction of the SDSS DR6

Wiener-filter with the ARGO code: FK, Jasche, Li, Ensslin, Metcalf, Wandelt, Lemson & White 2009

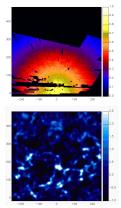


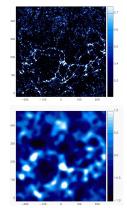
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Results: detection of a super-void in the SDSS DR6

(about 250.000 galaxies from the main sample)





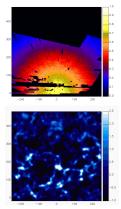
FK, Jasche, Li, Ensslin, Metcalf, Wandelt, Lemson & White 2009

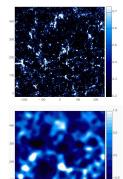
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Results: detection of a super-void in the SDSS DR6

(about 250.000 galaxies from the main sample) \rightarrow cluster prediction



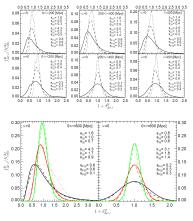


FK, Jasche, Li, Ensslin, Metcalf, Wandelt, Lemson & White 2009

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Results: matter statistics in the SDSS DR6



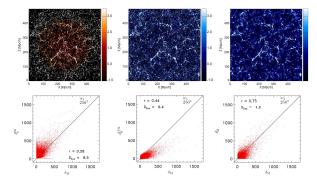
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Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Results: lognormal filter against Wiener filter and inverse weighting

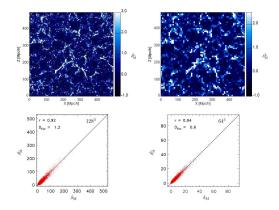
tests with the Millenium Run including selection function effects (about 350.000 mock galaxies)



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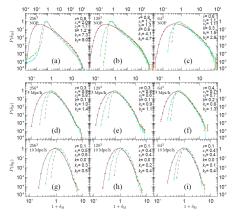
Results: lognormal filter against Wiener filter and inverse weighting



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Results: matter statistics in the lognormal reconstruction



FK, Jasche & Metcalf 2009 (upgrade of the ARGO code)

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Sampling the posterior

 Hamiltonian sampling (Taylor et al 2010, Jasche & FK 2010, FK, Simona Gallerani & Andrea Ferrara 2010)

$$H(\mathbf{s},\mathbf{p}) = K(\mathbf{p}) + E(\mathbf{s}),$$
 (22)

kinetic term with a given mass as the variance for the momenta

$$\mathcal{K}(\mathbf{p}) = \frac{1}{2} \mathbf{p}^{\dagger} \mathbf{M}^{-1} \mathbf{p}, \qquad (23)$$

Marginalization over the momenta

$$P(\mathbf{s},\mathbf{p}) = \frac{e^{-H}}{Z_H} = \frac{e^{-K}}{Z_K} \frac{e^{-E}}{Z_E} = P(\mathbf{p})P(\mathbf{s}),$$
(24)

Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Sampling the posterior

• Hamiltonian evolution equations: $(\mathbf{s}, \mathbf{p}) \rightarrow (\mathbf{s}', \mathbf{p}')$

$$\frac{\mathrm{d}\mathbf{p}}{\mathrm{d}t} = -\frac{\partial H}{\partial \mathbf{s}} = -\frac{\partial E}{\partial \mathbf{s}},$$
(25)
$$\frac{\mathrm{d}\mathbf{s}}{\mathrm{d}t} = \frac{\partial H}{\partial \mathbf{p}} = \mathbf{M}^{-1}\mathbf{p},$$
(26)

Metropolis-Hastings acceptance step

$$p_a = \min(1, e^{-\delta H}), \tag{27}$$

 $\delta H = H(\mathbf{s}', \mathbf{p}') - H(\mathbf{s}, \mathbf{p}) \rightarrow$ we do not care about the evidence!

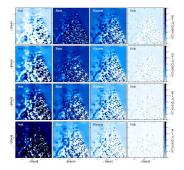
Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Results: matter field reconstruction of the SDSS DR7 with Hamiltonian sampling and the lognormal prior

deformation tensor of the grav.
 Pot. Φ:

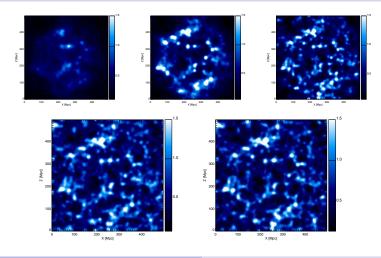
$$T_{ij} = \frac{\partial^2 \Phi}{\partial x_i \partial x_j} \qquad (28)$$

- eigenvalues $\lambda_1 \ge \lambda_2 \ge \lambda_3$, > 0: contraction, < 0: expansion
- classification: void: all < 0, sheet $1\lambda > 0$, filament $2\lambda > 0$, halo: $3\lambda > 0$ (with threshold see Forero-Romero et al 2009) Jasche, FK, Li & Ensslin 2010



Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

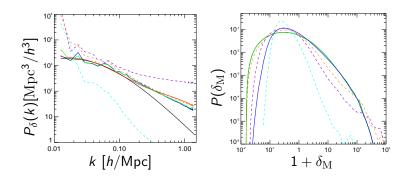
Skewed matter statistics: FK in prep



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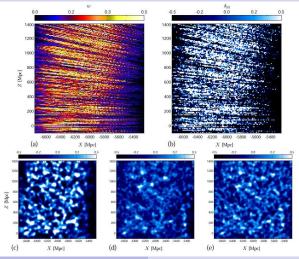
Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Skewed matter statistics



Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

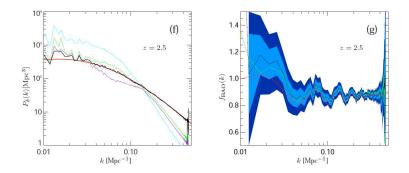
Lyman alpha forest 3D



Francisco-Shu Kitaura

Definition of the prior Definition of the likelihood Link between the prior and the likelihood Bayes theorem: the posterior Maximum a posteriori MCMC Sampling

Lyman alpha forest 3D



FK, Simona Gallerani & Andrea Ferrara 2010 Simona Gallerani, FK & Andrea Ferrara 2010

Conclusions

- There is a need to compare observations with theory as precisely as possible.
- Observations are plagued by many uncertainties which require a statistical treatment.
- The Bayesian approach is flexible and clear.
- We have shown that we can deal with complex models in this framework.
- There is a lot to do in Cosmology!